A STUDY OF EXPENDITURE PATTERNS OF CALCUTTA MIDDLE CLASS FAMILY BUDGETS

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The traditional method in family expenditure studies is to start with some classification based on traits fixed arbitrarily from our knowledge of socio-economic conditions such as community, province or some arbitrary economic levels. From such classified data according to immediately noticeable traits attempts are male to calculate average expenditure patterns of each group and to explain the variation of the expenditure patterns due to the basic classification itself or due to enuses indirectly involved in the classification itself. In the middle-class budget survey conducted in 1945 by the Indian Statistical Institute data was collected from 1785 family schedules in three independent sub-amples. Analysis of the data pertaining to one of the three sub-amples on the basis of a classification of families according to total family expenditure reveals that the family expenditure patterns do not appreciably differ from group to group. Table I giving expenditure patterns compiled from data collected from Calcutta diet survey (1943) reveals the truth of the above statement.

Table 1. Expenditure patterns. expenditure groups in rupees.

51- 101- 151- 201- 251- 301- 401- cons

							2.71	401-	6003
	3u ·	[196]	150	54143	250	300	41K)	600	above
essential fond items	BO -02	47 -541	41.72	39 -34	34 · 30	31 -72	71 -57	27 - 17	:3-74
non-executial food items	9 -00	17 -40	20 60	21 -bst	23 -24	22 - 31	26 - 14	27 -96	28 49
rent & fuel	18 -29	145 - 4.5	10 -45	10.21	14 -51	14:51	13.52	13 - 15	9.51
miscellaneous & clothing	13.40	18 -59	21.17	23 - 35	25 Pl	5× 44	28 - 77	31 - 71	34 -05
total	100-00	HHO -(H)	100 -00	100 -00	PHO -00	100 -00	1(n) (p)	100 40	100.00
average family size	3 - 27	4-11	4 -83	7 - 14	A -22	8 . 174	10 -36	11.77	16 -12
Though the figures sh	ow that	the pat	terns ec	onsistent	ly vary	with	total e	xpendi	ture an
examination of the in	տնւմներու	family	variati	on reven	Is that	only a	small fr	action	of it is
ascribable to total ex									
family composition, co									
the variation of expenditure patterns. The average family size increases consistently with									
the expenditure level as shown by the same table. In fact one may state that the									
classification adopted here does not separate the families into distinct types with respect									
to their expenditure patterns. Hence any conclusions based on the above type of analysis									
may not reveal the actual state of affairs. In fact cost of living indices for 1945 for all the									
groups calculated with 1939 as the base year were found to lie between 280 and 290 though									
one should expect the cost of living index should be high for high expenditure levels where									
proportionately more is spent on items (like fish, meat etc.) the prices of which have arisen									
by 5 or 6 times that o	f the ure	WAT DE	ice.						

Therefore, an attempt has been made in this paper to classify families according to expenditure pattern, and then to investigate further in what respects families possessing the same pattern agree between themselves and in what manner they contrast with families possessing a different pattern. This method of tackling the problem may help us to clussify families into distinct types by means of some common traits. If no such common traits are found as may happen sometimes when manifold classifications are required to explain the variations of expenditure pattern then we may classify the families according to expenditure

percentage of

23

expenditure on

pattern itself and study the effect of several factors on each group separately at different times. This is a better method of study of the economic conditions of families at different intervals of time than the one usually adopted which will give a distorted picture of the existing conditions.

The expenditure pattern studied here consists of sets of four variates which are (a) proportion of expenditure on essential food items (by essential food items is meant cereals, pulses, vegetables, spices and gur) to total expenditure (b) proportion of expenditure on nonessential food items (all food items other than included in (a)) to total expenditure (c) proportion of expenditure on rent and fuel to total expenditure (d) proportion of expenditure on miscell meous and clothing to total expenditure. This effects a classification of all items into four homogeneous classes with price relatives of items in each class close together. As all the items of expenditures are covered up by (a), (b), (c), (d) it is not necessary to echsider the four variates as they will add up to one. It is sufficient, therefore, to deal with any three of them only in our statistical analysis.

Since the statistical analysis with a multiplicity of variates is complicated we shall try to search for a function of these variates which can explain a major portion of the variance of the sets of variates. This function we shall call as the most discriminating function and we shall replace the original set of three variates which represented the family pattern by this function.

Each family can be represented by a point in three dimensions with coordinates p1, p2, p3 representing the proportional expenditures on (a), (b) and (c). The variance of such points can be resolved into three components by taking any three mutually orthogonal axes. We shall choose that system of axes for which the sum of squares of the projections of the point vectors on each of the axes is stationary. If the variance of the projections on an axis constitutes a major component of the total variance then the original points may be replaced by their projections on this line. Having obtained the equation of the line we shall get the distance of each projected point from the origin and thus replace the set of variates (p1, p2, p3) by this one variate, namely the distance of the projected point on this line from the origin.

The mathematical derivation of the linear components described above is given below. Let $l_1 p_1 + l_2 p_3 + l_3 p_3 = t$ be one of the components required, then variance of t is

$$\sum_{j=1}^{n} \sum_{i=1}^{n} l_i l_j v_{ij} = Q$$
 (1)

where v_{ij} is the variance of p_i and v_{ij} is the covariance between p_i and p_j . Also since the sum of the variances of the three components is equal to the total variances of the set py, pa, pa we must have $l_1^2 + l_2^2 + l_3^2 - 1 = v = 0$.. (2) So then our object is to maximise Q with respect to l1, l2 and l3 subject to the restriction (2). Using a Lagrangian multiplier we obtain l1, l2 and l3 as solutions of (2) and (3)

$$\frac{\partial Q}{\partial l_i} - k \frac{\partial v}{\partial l_i} = 0$$
 $i = 1, 2, 3$ (3)

On simplifying (3) we get

$$\begin{cases} (V_{11} - k)l_1 + V_{12}l_2 + V_{13}l_3 = 0 \\ V_{21}l_1 + (V_{22} - k)l_1 + V_{22}l_3 = 0 \\ V_{21}l_1 + V_{22}l_2 + (V_{23} - k)l_3 = 0 \end{cases}$$
 (4)

If solutions other than $l_1=l_2=l_2=0$ exist the determinant of the set of equations (4) must wanish. This gives

$$\begin{bmatrix} \begin{pmatrix} V_{13}-k \end{pmatrix} & V_{13} & V_{13} \\ V_{11} & (V_{12}-k) & V_{13} \\ V_{21} & V_{42} & (V_{22}-k) \end{bmatrix} = 0$$
(5)

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which is a cubic equation in k called the characteristic equation and it is a well known result that such equations have all the roots real and positive. Let k_1 , k_1 and k_2 represent the three roots of $\{5\}$. Substituting any one of them in (4) and multiplying 1st member of (4) by l_1 , and 3rd by l_2 and adding up we get $Q = k(l_1^2 + l_2^2 + l_3^2) = 0$ i.e. Q = k. Hence each of these roots stand for the variance of each one of the components possessing the properties described above. Substituting each one of the roots k_1 , k_2 and k_2 successively in the first two members of (4) and using (2) we determine the corresponding l's by solving in the usual manner. Let us denote the first, accord, and third components thus obtained as $l_1 l_2 + l_4 l_2 + l_4 l_3 p_3$, i = 1, i =

One of the three subsamples collected in Calcutta diet survey consisting of 625 family schedules is taken up for analysis. Since 128 family schedules belonged to house owning families whose housing expenses are naturally very low were omitted in our analysis and the results of the analysis on the remaining 497 family schedules are given below. Division of the sums of squares or sums of products by total sample size is avoided to facilitate computational work. The characteristic equation derived from the covariance matrix is

(8·4554-k) -4·2233	-4-2233 (2-7997-k)	-C963 -1-0383	=0
-0963	-1-0383	(1·4905-k)	
L3-12-7267k3-	+21.5247k + .4018	3=0	

which is same as

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The roots of this equation are $k_1 = 10$ -7461, $k_2 = 1$ -9807, $k_3 = 0$ -0189. Total variance sequal to 10-7461+1-9807+0-0182=12-7267. The contribution by the first component to the total variance of the set is 814% and that by other two components tegether is only 15-6% of the total variance. Hence replacing the family pattern (p_1, p_2, p_3) by a single component as the first one we are leaving only about 10% of the total variation of the expenditure pattern unaccounted and this cannot vitinte the results of our statistical analysis. The three components as well as their variance contributions are given below,

mponenta	direction	direction cosmes of the axes					
	l_{i}	l_2	l,				
1	·8780	-4748	-0621	84-4%			
2	-2933	4302	8537	15.5%			
3	-3783	-7678	-5170	0.19/			

Also it can easily be verified that these components are orthogonal to each other.

Taking the first component which explains 84-4% of the total variance to represent the family pattern we find that with increase of family size this component increases but with increase of total expenditure it decreases as is shown by Table 2 giving mean values of the principal component of expenditure pattern by total expenditure groups and family sizes,

Table 2. Variation of X (principal component) with family size and total expenditure.

family size	exponditure group in rupers							
	0-50	51-100	101-150	151-200	201-250	231-300	301-400	401-400
2	141	-30	-10	-02	_	-01	_	10.
3	-52	-30	-16	-04	-06	_		_
4	-60	-31	-23	-15	-20	-12		-06
5	_	•39	-29	-21	-26	-14	-03	

Except in the case of the groups 151-200 and 201-250 in family sizes 4 and 5 we faul that variations are consistent with the statement made above. Such discrepancies as observed are due to a samples being too small in those groups. Let us next study the variation in expenditure pattern due to total expenditure by family sizes. Table 3 below gives the regression equation of x, the principal component on t, the total expenditure for each family size and the contribution of variance due to total expenditure.

Table 3. Regression of the principal component of expenditure

family Fize	no. of families	regression equation	variance of x	variance of z due to t	percentage variation due to t	correlation between #
(1)	(2)	(3)	(4)	(5)	(6)	(7)
1.3	34	$(x - \cdot 2003) = - \cdot 0013(t - 93 \cdot 59)$	-0480	-3715	39 - 15	- 0257
3	43	$(x - \cdot 2723) = - \cdot 0029(t - 101 \cdot 67)$	1 -4133	-9732	68.80	8311
ĭ	70	$(x - \cdot 2602) = - \cdot 0011(t - 130 \cdot 94)$	1-3842	-2012	26 -23	6019
5	72	$(x - \cdot 2992) = - \cdot 0015(t - 131 \cdot 93)$	1.0546	-5575	52 86	7271
6	60	$(x - \cdot 2778) = - \cdot (x018(t-157 \cdot 06))$	1 -2260	-8084	62 -80	-·8117
7	46	$(x - \cdot 2559) = - \cdot 0010(t - 189 \cdot 99)$	-6974	-2867	41 -11	6412
8	46	$(x-\cdot 2417) = -\cdot 0012(t-220\cdot 47)$	-4311	-2490	57 - 76	7600
9	3.5	$(x - \cdot 2090) = - \cdot 0004(t - 271 \cdot 33)$	1 -2219	-2169	17 - 75	4213
10	32	(z2418) =0002(t-323.72)	-6330	.1057	16.70	~· t089

It can be seen that even for a given family size the variation brought about by total expenditure is not very appreciable except in the case of families of sizes 3, 5, 6 and 8. If we had analyzed the variance irrespective of family size the variance contribution by total expenditure would have been very much smaller as is shown by the following analysis.

Combining families of all sizes we find the correlation between total expenditure and the principal component reduces to -3290 and the variance contributed by total expenditure to the total variance of the pattern is only 10-82%. This explains why, when classified according to total family expenditure, average family patterns showed little variation between expenditure groups.

Similar analysis has not been done on families whose size execeds 10 since samples are very small in those cases. The variation from one family size to another cannot be given a functional form (as in the above case we had given a linear fit) since families of a given size belong to a special type distinct from others with respect to composition. As family composition determines to a very larger extent the variation in pattern the variation due to size gets distorted and hence the variation due to family size and total expenditure have not been estimated by applying partial regression formulae to the data.

SUMMARY,

- (1) Since the principal component $l_1, p_1 + l_{12}, p_2 + l_{13}, p_3$ explains a major portion of the variance, there exists a single economic factor which influences to a great extent the expenditure pattern. This component which increases with family size and decreases with total expenditure can be used to classify families according to poverty levels for budgetary studies.
- (2) Though family size and total expenditure together explain in many cases a large part of the variation in expenditure patterns we find that there exist other factors like family composition etc., which too in no small measure determine the poverty level mentioned abore.
- (3) Since classification by expenditure levels alone cannot divide the families into distinct types with respect to expenditure pattern, any study of changes of pattern of families in a given expenditure level after the lapse of a period leads to erroneous judgement. Therefore, it would be better to group the families according to poverty levels determined by the principal ebmponent derived in this article and study how they change their groups and assess improvement or deterioration.